An Unbiased Analysis of the Proton's Elastic Form Factors

GE & GM



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The Form Factors

F(q) = ?

Proton current is parametrized by general form factors

$$J^{\mu}_{\gamma} = \bar{u}_{N}(p') \Gamma^{\mu}_{\gamma}(q) u_{N}(p)$$

- Only asymptotically constrained by theory
 - Need Experimental data to understand further

The Form Factors

F(q) = ?

Proton current is parametrized by general form factors

$$J^{\mu}_{\gamma} = \bar{u}_{N}(p') \Gamma^{\mu}_{\gamma}(q) u_{N}(p)$$

$$\Gamma_{\gamma}^{\mu}(q) = \gamma^{\mu} F_1 \left(Q^2 \right) + \frac{i \sigma^{\mu \nu} q_{\nu}}{2M} F_2 \left(Q^2 \right)$$

- Only asymptotically constrained by theory
 - Need Experimental data to understand further

Observables

<u>dσ</u> dΩ

 Connect observables to Form Factors

$$G_E(Q^2) = F_1(Q^2) - \tau F_2(Q^2)$$

$$G_M(Q^2) = F_1(Q^2) + F_2(Q^2)$$

• LT: OPE Cross Section:

• PT: Polarized Cross Sections:

Observables

$\frac{\mathsf{d}\sigma}{\mathsf{d}\Omega}$

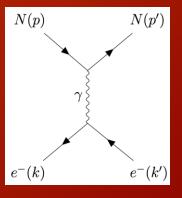
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• PT: Polarized Cross Sections:

Observables

$\frac{d\sigma}{d\Omega}$

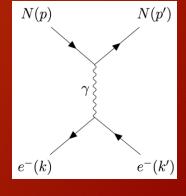
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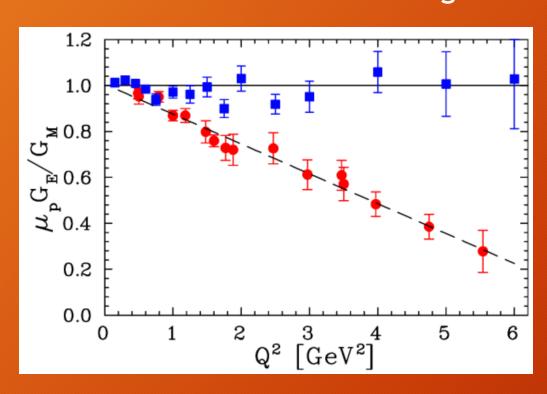
$$rac{d\sigma^{(T)}}{d\Omega} \propto ~G_E G_M$$

$$rac{d\sigma^{(L)}}{d\Omega} \propto \, G_M^2$$

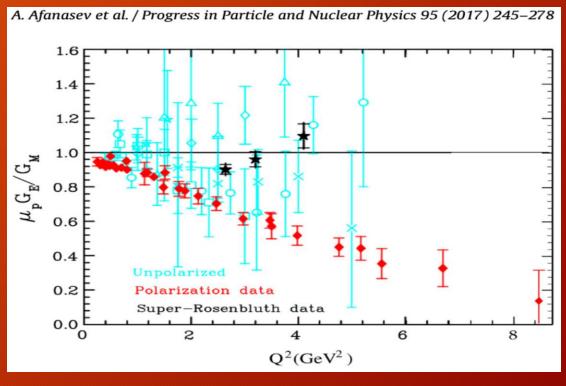
Two Independent Form Factor Ratios

LT ≠ PT

Significant Disagreement!



Single Experiment

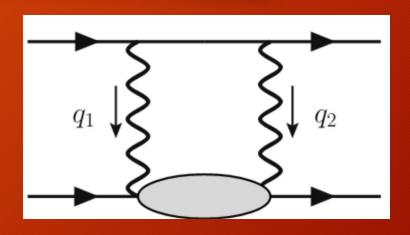


Several Experiments

Probable Causes of Discrepancy

T ≠ PT

Two Photon Exchange Corrections



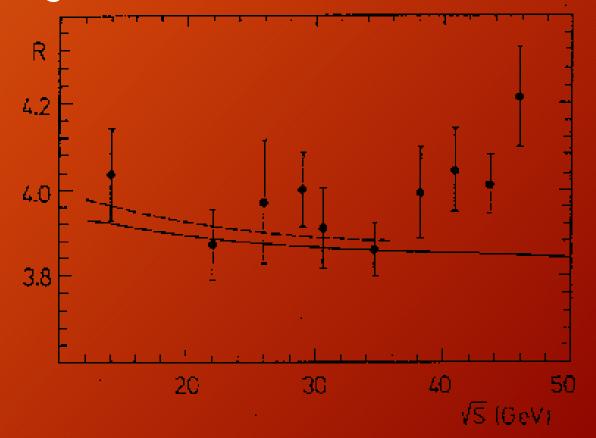
- Multiplicative Uncertainty
 - Correlates Whole Experiment

$$y_i \cdot (1 \pm \Delta n_i)$$

Multiplicative Uncertainty - How?

- Improper treatment leads misleading fits
 - "Peelle's Pertinent Puzzle" @ Cello

$$y_i \cdot (1 \pm \Delta n_i)$$



Traditional Fitting and The Penalty Trick

• Chi-square comes from Gaussian
$$P(y_1,y_2,\ldots,y_N|M_i)=\prod_{i=1}^N \frac{1}{\sqrt{2\pi}\Delta y_i}e^{-\frac{1}{2}(y_i-M_i)^2/(\Delta y_i)^2}$$

$$\chi^2(oldsymbol{lpha}) = \sum_{i=1}^N rac{(y_i - M_i)^2}{\Delta y_i^2}.$$

- Penalty Trick

• Scaling Factors Biased
$$\sum_{i=1}^{\mathcal{N}} \left[\frac{(n_i-1)^2}{(\Delta n_i)^2} + \sum_{j=1}^{N_i} \frac{(n_i \cdot y_{ij} - M_{ij})^2}{(\Delta y_{ij})^2} \right]$$

Blueprint — The t₀ Method

$$y_i \cdot (1 \pm \Delta n_i)$$

Blueprint — The to Method

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$$\sum_{i=1}^{\mathcal{N}} \left[\sum_{j=1}^{N_i} rac{(y_{ij} - \mathrm{M}_{ij})^2}{(\Delta y_{ij})^2 + (y_{ij}\Delta n_i)^2}
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Blueprint — The t₀ Method

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$$\sum_{i=1}^{N} \left[\sum_{j=1}^{N_i} \frac{(y_{ij} - M_{ij})^2}{(\Delta y_{ij})^2 + (M_{ij} \Delta n_i)^2} \right] \longrightarrow \sum_{i=1}^{N} \left[\sum_{j=1}^{N_i} \frac{(y_{ij} - M_{ij})^2}{(\Delta y_{ij})^2 + (\hat{M}_{ij} \Delta n_i)^2} \right]$$

Blueprint — The t₀ Method

• 1) Aforementioned iterative guess

- 2) Monte-Carlo replica averaging
 - Best model is average of replica best fits (non-linearity causes issues)

$$F_{\text{best}}(Q^2; \alpha) = F_{\text{avg}}^i$$

Extending the to Method — MOP Covariance Matrix

MOP

Covariance Matrix Ambiguous

Model Outer Product

$$\begin{pmatrix} (\Delta y_{i1})^2 & 0 \\ 0 & (\Delta y_{i2})^2 \end{pmatrix} + \begin{pmatrix} \hat{\mathbf{M}}_{ij} \Delta n_i \end{pmatrix}^2 \\ \begin{pmatrix} (\Delta y_{i1})^2 + \begin{pmatrix} \hat{\mathbf{M}}_{i1} \Delta n_i \end{pmatrix}^2 & \begin{pmatrix} \hat{\mathbf{M}}_{i?} \Delta n_i \end{pmatrix}^2 \\ \begin{pmatrix} \hat{\mathbf{M}}_{i?} \Delta n_i \end{pmatrix}^2 & (\Delta y_{i2})^2 + \begin{pmatrix} \hat{\mathbf{M}}_{i2} \Delta n_i \end{pmatrix}^2 \end{pmatrix}$$

Extending the to Method — MOP Covariance Matrix

- Covariance Matrix Ambiguous
- Model Outer Product

$$\begin{pmatrix} (\Delta y_{i1})^2 & 0 \\ 0 & (\Delta y_{i2})^2 \end{pmatrix} + \begin{pmatrix} \hat{\mathbf{M}}_{ij} \Delta n_i \end{pmatrix}^2 \\ \begin{pmatrix} (\Delta y_{i1})^2 + \begin{pmatrix} \hat{\mathbf{M}}_{i1} \Delta n_i \end{pmatrix}^2 & \begin{pmatrix} \hat{\mathbf{M}}_{i?} \Delta n_i \end{pmatrix}^2 \\ \begin{pmatrix} \hat{\mathbf{M}}_{i?} \Delta n_i \end{pmatrix}^2 & (\Delta y_{i2})^2 + \begin{pmatrix} \hat{\mathbf{M}}_{i2} \Delta n_i \end{pmatrix}^2 \end{pmatrix}$$

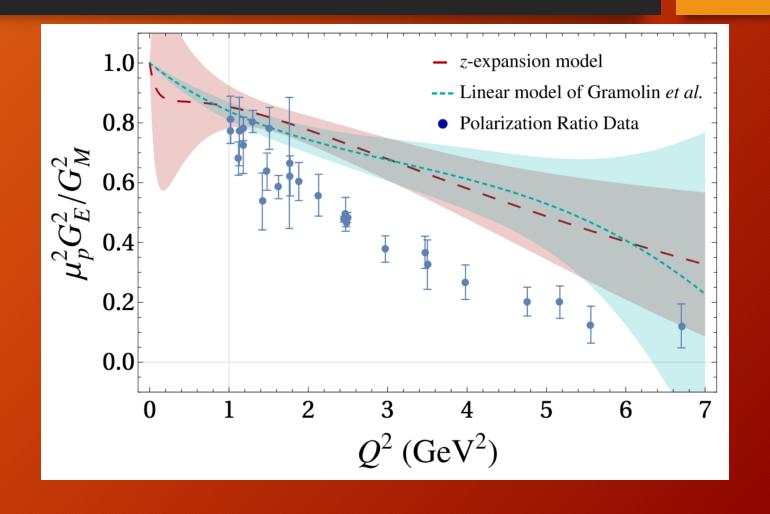
$$\begin{pmatrix} (\Delta y_{i1})^{2} + (\hat{M}_{i1}\Delta n_{i})^{2} & \hat{M}_{i1}\hat{M}_{i2} (\Delta n_{i})^{2} \\ \hat{M}_{i1}\hat{M}_{i2} (\Delta n_{i})^{2} & (\Delta y_{i2})^{2} + (\hat{M}_{i2}\Delta n_{i})^{2} \end{pmatrix}$$

n_i vs. t₀

Penalty Trick vs. MOP Method

- Very Similar Results
 - Penalty Trick is still a good estimator

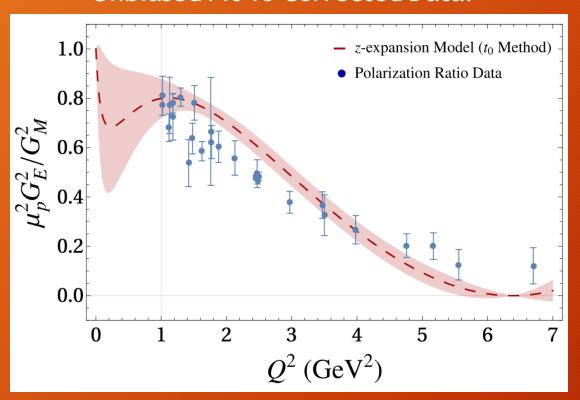
- Main takeaways
 - LT still not equal to PT
 - Fitted normalizations are merely a crutch



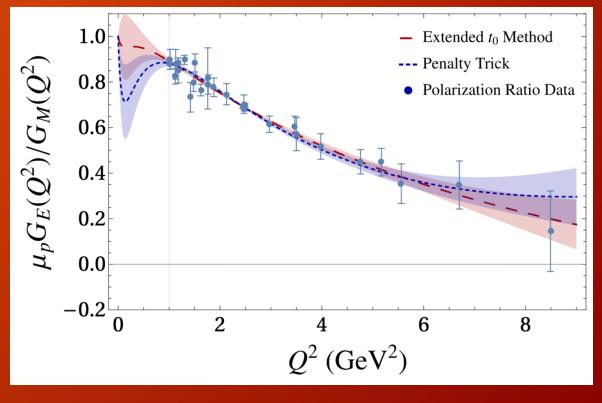
Form Factor Ratio Discrepancy Still At Large

LT ≠ PT

- Need Two-Photon-Exchange Corrections
 - Unbiased Fit To Corrected Data:



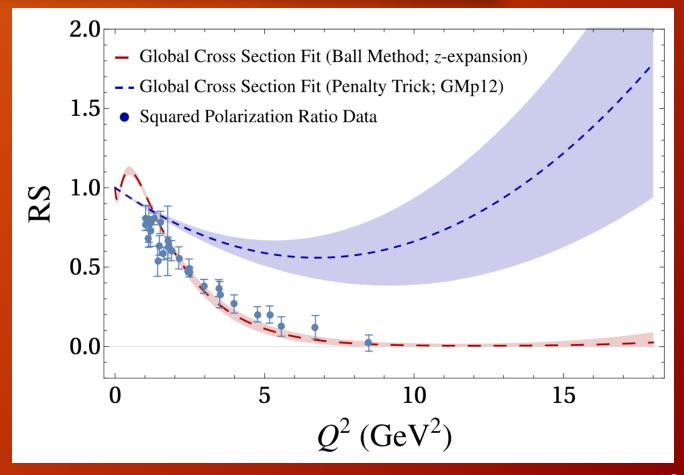
14.5% increase chi-square



Improvements can be made to TPE

• If one treats Normalization Error as point-to-point error:

• Is multiplicative error grossly overestimated?

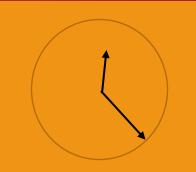


Summary

:)

- New method for unbiased fitting of non-linear models
- Current TPE Corrections help significantly close gap, room for improvement
- Scale Uncertainties are likely being over-estimated
- Normalization Factors are Merely a Crutch
- Future Work
 - Perform updated Low-Q² data as in Bernaur (2014)
 - Does the updated fitting procedure effect proton radius?

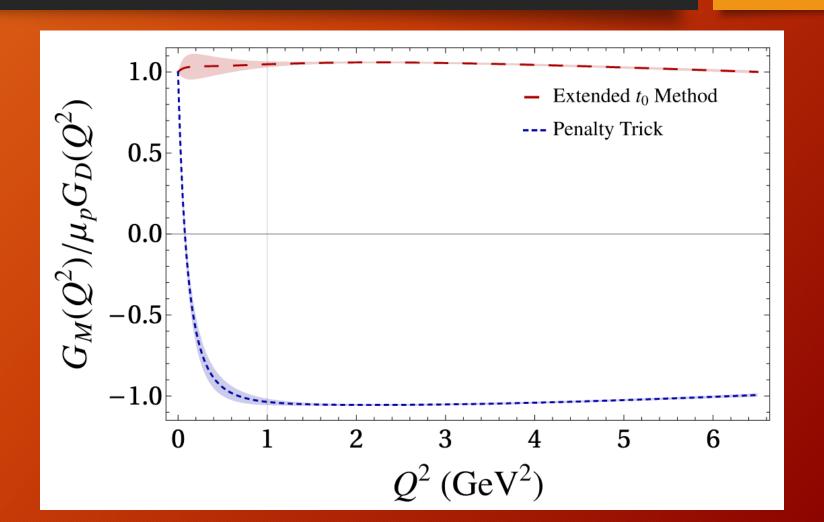
Question Time



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Perils of Non-linear fitting

 Non-linear fits need a lot of supervision



Extending the to Method — Non-Linear Models

- Iterative parameter search only good for Linear models
 - Only when model is linear: Average of models is the model of averaged parameters
- Non-Linear Models
 - Use average parameters and hope for convergence (works surprisingly well)
 - If necessary, can use L² norm to find 'closest' model to average model

$$\int_{x_{\min}}^{x_{\max}} \left(F\left(x, \boldsymbol{\alpha}\right) - \bar{f}_i \right)^2 dx$$

Multi-Experiment Rosenbluth Extraction

 Without considering full covariance matrix Rosenbluth Extractions are not useful

